

Part (A): Review of Probability [Statistics I revision]

1 Definition of Probability

1.1 Experiment

An experiment is any procedure whose outcome is uncertain

- toss a coin
- throw a die
- buy a lottery ticket
- measure an individual's height h

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1.2 Sample Space

A sample space, S , is the set of all possible outcomes of a random experiment

- $S = \{H, T\}$
- $S = \{1, 2, 3, 4, 5, 6\}$
- $S = \{WIN, LOSE\}$
- $S = \{h : h \geq 0\}$

1.3 Events

An event, A , is an element, or appropriate subset of S

e.g. roll a die, $S = \{1, 2, 3, 4, 5, 6\}$

- event A , that an even number is obtained is given by $A = \{2, 4, 6\}$

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- event B , that a number no greater than 4 is obtained is described by $B = \{1, 2, 3, 4\}$

S can be **discrete** (e.g. coin, die, lottery), or **continuous** (e.g. height).

In the first part of this course we will deal with the discrete case (as in Stats 1).

Set operations

If A, B are events, according to Set Theory, we define the operations:

- $A \cup B$: union, 'A or B happening'
- $A \cap B$: intersection, 'A and B happening'
- A^c (or A', \bar{A}): complement of A wrt S , 'not A'
Notice that: $A \cup A^c = S, \quad A \cap A^c = \emptyset$

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1.4 Probability mass function

A probability function, $P(s)$, is a real-valued function of a collection of events from a sample space, S , attaching a value in $[0, 1]$ to each event, satisfying the **axioms**:

A1. $P(A) \geq 0$ for any event A

A2. $P(S) = 1$

A3. If A_1, A_2, A_3, \dots is a countable collection of mutually exclusive events ($A_i \cap A_j = \emptyset$ for $i \neq j$, i.e. they have no common elements), then $P(\cup_i A_i) = \sum_i P(A_i)$

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Important properties of $P(s)$

- $P(\emptyset) = 0$
- $P(A \cup B) = P(A) + P(B) - P(A \cap B)$
This can be generalised for n events, e.g. for $n = 3$:
 $P(A_1 \cup A_2 \cup A_3) = P(A_1) + P(A_2) + P(A_3) - P(A_1 \cap A_2) - P(A_1 \cap A_3) - P(A_2 \cap A_3) + P(A_1 \cap A_2 \cap A_3)$.
Notice: $P(A \cup B) = P(A) + P(B)$ **not always true!**
It only holds for disjoint (mutually exclusive) events.
- $P(A^c) = 1 - P(A)$
Proof. A, A^c are disjoint, and $A \cup A^c = S$
 $\Rightarrow P(A \cup A^c) = P(S) = 1$
 $\Rightarrow P(A) + P(A^c) = 1$
In many practical situations it is easier to calculate $P(A^c)$ first and use this property to calculate $P(A)$.

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1.5 Independence and Conditional Probability

For any event A with $P(A) > 0$ we can define

$$P(B/A) = \frac{P(A \cap B)}{P(A)} \quad (1)$$

If A and B are **independent** then by definition $P(A \cap B) = P(A) \times P(B)$ and therefore

$$P(B/A) = \frac{P(A)P(B)}{P(A)} = P(B).$$

Note that rearrangement of (1) gives the chain (multiplication) rule:

$$P(A \cap B) = P(A)P(B/A) = P(B)P(A/B). \quad (2)$$

This can also be generalised for n events.

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1.6 Total Probability Rule

If A_1, A_2, \dots, A_k are mutually exclusive events and form a partition of a sample space S (i.e. $\cup_{i=1}^k A_i = S$), with $P(A_i) > 0 \quad \forall i$,

then for an event $B \in S$:

$$P(B) = \sum_{i=1}^k P(B \cap A_i) = \sum_{i=1}^k P(A_i)P(B/A_i) \quad (3)$$

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Example:

An insurance company covers claims from 4 different motor portfolios, A_1, A_2, A_3 and A_4 . Portfolio A_1 covers 4000 policy holders; A_2 covers 7000; A_3 covers 13000; and A_4 covers 6000 of the total 30000 policy holders insured by the company. It is estimated that the proportions of policies that will result in a claim in the following year in each of the portfolios are 8%, 5%, 2% and 4% respectively. What is the probability that a policy chosen randomly from one of the portfolios will result in a claim in the following year?

Solution ...

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1.7 Bayes Theorem

If $P(A) > 0$ and $P(B) > 0$, from (1)

$$P(A/B) = \frac{P(A \cap B)}{P(B)} = \frac{P(A)P(B/A)}{P(B)} \quad (4)$$

or, if A_1, A_2, \dots, A_k form a partition of a sample space S with $P(A_i) > 0 \quad \forall i$, and $P(B) > 0$, then from (3)

$$P(A_j/B) = \frac{P(A_j)P(B/A_j)}{\sum_{i=1}^k P(A_i)P(B/A_i)}, \quad j = 1, 2, \dots, k. \quad (5)$$

Example: (previous cont.)

If a claim rises from a policy in the year concerned, what is the probability that the policy belongs to portfolio A_3 ?

Solution ...

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1.8 Random variables

Until now we have associated probabilities with events in a sample space. We will now review the concept of random variables.

Definition:

A random variable (r.v.) X is a function from a sample space S to \mathbb{R} (the real numbers).

That is, to any outcome, s , we associate a real number $X(s)$.

The **range** of a r.v. X , denoted as S_x , is the set

$\{r \in \mathbb{R} : r = X(s) \text{ for some } s \in S\}$.

[i.e. the set of all real numbers r which are equal to $X(s)$ for some outcome s .]

We will usually denote random variables using capital letters, e.g. X , and their realisations (values) using small letters, e.g. x .

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If S is **finite** or **countable** then the range of X is a set of discrete numbers. We say X is a **discrete random variable**.

Some simple examples:

(i) Experiment: Toss a balanced coin 3 times.

Sample space S : All possible sequences of T, H of length 3.

Let $X(s)$ = no. of Heads in outcome (sequence) s . Then

X is a discrete random variable.

(ii) Experiment: Toss a balanced coin until a 'H' is thrown.

Sample space: $S = \{H, TH, TTH, \dots\}$.

Let $X(s)$ = no. of throws (i.e. length of s). Then

X is a discrete random variable.

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1.9 Probability function of a discrete r.v.

Any probability function on S leads naturally to a probability function on the range of X , defined by

$$f_X(x) = P(X = x) = \sum_{s: X(s)=x} P(s).$$

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In **Example (i)** before:

$s :$	HHH	HHT	HTH	THH	TTT	TTH	THT	HTT
Prob:	$\frac{1}{8}$							
$X :$	3	2	2	2	0	1	1	1

Probability of 2 H s is

$$P(X = 2) = \frac{1}{8} + \frac{1}{8} + \frac{1}{8} = \frac{3}{8}.$$

Usually we consider r.v.'s which describe some useful summary of the outcome of an experiment, e.g. the total score in a multiple choice test.

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Axioms of probability

For a probability function of a discrete r.v. to satisfy the axioms of probability, it is sufficient that

- i. $f_X(x) \geq 0$ for all $x \in S_x$
- ii. $\sum_{\text{all } x} f_X(x) = 1$

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1.10 S infinite

Until now in the examples we have only considered cases where the sample space S consists of a countable and finite number of events. However, there may be also an **infinite number of possibilities** involved with the experiment under consideration.

Some results on infinite series ...

Example:

Two (TV show) players (A and B) play a game in which each wins a prize if s/he is the first to choose the correct screen out of the 6-screen display panel shown to them. Player A goes first, and the prize is randomly re-allocated to one of the screens after each player's choice (so that the probability of either player winning at a single draw is independently $1/6$). What is the probability that A wins?

Solution ...

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2 Some common discrete distributions [Statistics I]

2.1 The binomial distribution: $\text{Bin}(n, p)$

Consider an experiment consisting of n **independent** trials where the probability of success is p , $0 < p < 1$.

Let X denote the number of successes (S's). Then

$$f_X(x) = P(X = x) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, 1, 2, \dots, n. \quad (6)$$

Why? The sample space S consists of all sequences of S and F of length n . For any such sequence, s , with x S's and (therefore) $(n-x)$ F's the probability is $P(s) = p^x (1-p)^{n-x}$

(since the trials are independent so probabilities multiply).

There are $\binom{n}{x}$ such sequences, and therefore (6) follows.

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2.2 The geometric distribution: $\text{Geometric}(p)$ [Statistics I revision]

Again consider a sequence of independent trials with probability of 'success' $P(S) = p$. Let X be the number of trials required to achieve the 1st success [e.g. number of attempts required to pass the driving test].

$S = S, FS, FFS, FFFS, \dots$

$X = 1, 2, 3, 4, \dots$

Therefore for $x = 1, 2, 3, \dots$

$$f_X(x) = P(X = x) = P(\overbrace{FF \dots F}^{x-1} S) = (1 - p)^{x-1} p. \quad (7)$$

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The memoryless property

The geometric distribution has the memoryless property. That is, given that there have already been n trials without success, the probability that x additional trials will be required for the first success is independent of n :

$$P(X > x + n / X > n) = P(X > x) \quad \text{[Show]}$$

[Is the number of attempts required to pass the driving test well described by the geometric distribution?]

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2.3 [NEW!] The negative binomial distribution: $\text{NBin}(r, p)$

This is a **generalisation of the geometric distribution**.

Consider the same sequence of independent trials as in (2.2) and let X denote the number of trials required to achieve r successes, where r can be any positive integer. Clearly the range of X is $r, r + 1, r + 2, \dots$

What is $f_X(x) = P(X = x)$, $x \geq r$?

Let s be an outcome for which $X(s) = x$. Then s is a sequence of S's and F's such that:

- (i) s has length x
- (ii) s has r S's and $x - r$ F's
- (iii) the last entry is S

From (ii), $P(s) = (1 - p)^{x-r} p^r$ (independence of trials!!).

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How many such outcomes are there?

Trial:	1	2	3	...	$x - 1$	x
Result:	?	?	?	...	?	S

$(r - 1)S's \ \& \ (x - r)F's$

There are $\binom{x-1}{r-1}$ ways of assigning the S's to position $1, 2, \dots, x - 1$.

Therefore

$$f_X(x) = P(X = x) = \binom{x-1}{r-1} (1 - p)^{x-r} p^r, \quad x = r, r + 1, \dots \quad (8)$$

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Relationship to geometric distribution

1. $\text{NBin}(1, p) \equiv \text{Geometric}(p)$
2. If X_1, X_2, \dots, X_r are independent $\text{Geometric}(p)$ r.v.'s, then $Y = X_1 + X_2 + \dots + X_r \sim \text{NBin}(r, p)$.

To summarise:

In a sequence of independent 'trials', each with probability of 'success' p :

- the number of successes in n trials follows a $\text{bin}(n, p)$ distribution
- the number of trials until the first success follows a $\text{geometric}(p)$ distribution
- the number of trials until the r th success follows a $\text{NBin}(r, p)$ distribution.

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2.4 [NEW!] The Poisson distribution: $\text{Poisson}(\lambda)$

This is a distribution which is often used to describe the outcomes of experiments that involve **counting objects or events** (e.g. the number of road accidents occurring on a stretch of road in a 1-month period).

Its range is $\{0, 1, 2, 3, \dots\}$, and its probability function is given by

$$f_X(x) = P(X = x) = \frac{e^{-\lambda} \lambda^x}{x!}, \quad x = 0, 1, 2, 3, \dots, \quad \lambda > 0. \quad (9)$$

It has connections with the $\text{Bin}(n, p)$ distribution, and also the Exponential distribution (see later).

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3 Expectation, variance and independence [Statistics I]

3.1 Mathematical Expectation of a discrete r.v. X

(Also referred to as **expected value** or **mean** of X .)

The expectation of a r.v. X is a weighted average of all possible values of X , with weights determined by the probability distribution of X . It measures where the centre of the distribution lies.

Definition:

Let X be a discrete r.v. with probability mass function $f_X(x)$. Then its expected value is

$$E(X) = \sum_{x \in S_x} x f_X(x) = \mu_x \quad (\text{mean of } X) \quad (10)$$

Notice that if the above sum does not converge for an infinite set S_x , then $E(X)$ is not defined.

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Properties of expectation:

If X, Y are r.v.'s and $a, b \in \mathbb{R}$ are constants:

- (i) $E(X + Y) = E(X) + E(Y)$
- (ii) $E(a) = a, \quad E(aX + b) = aE(X) + b$
- (iii) If $X \geq 0$, then $E(X) \geq 0$

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3.2 Expectation of a function of a discrete r.v.

Let X be a r.v. with probability function $f_X(x)$, and $Y = g(X)$ a real-valued function of it. Then Y is also a r.v., and its expectation is given by

$$E(Y) = \sum_{x \in S_x} g(x) f_X(x) = \sum_y y f_Y(y).$$

Example:

An individual invests an amount of £105k on a financial product. The (total) return of his investment at the end of the following year is given in the table:

Return (in £1000's) x :	118	113	110	107
Probability $f_X(x)$:	$\frac{3}{20}$	$\frac{7}{20}$	$\frac{7}{20}$	$\frac{3}{20}$

If the profit of the investment, Y , is given (as a function of the total return X) by the formula $Y = X^{0.99} - 105$, find the expected profit of the investor.

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Solution:

$$\begin{aligned} E(Y) &= \sum_x g(x) f_X(x) \\ &= (118^{0.99} - 105) \times \frac{3}{20} + (113^{0.99} - 105) \times \frac{7}{20} \\ &\quad + (110^{0.99} - 105) \times \frac{7}{20} + (107^{0.99} - 105) \times \frac{3}{20} \\ &= 7.5 \times \frac{3}{20} + 2.8 \times \frac{7}{20} - 0.1 \times \frac{7}{20} - 2.9 \times \frac{3}{20} = 1.6 \end{aligned}$$

[An expected profit of £1600. However, notice that with probability 1/2 the investor will lose money! Is this a 'good' investment?]

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3.3 Variance of a discrete r.v. X

The variance of a random variable x ($\text{var}(X)$, also σ_x^2) is defined as the expectation of the deviation (squared) of X from its mean (expected value), i.e.

$$\text{var}(X) = E\{X - E(X)\}^2 \quad (11)$$

and for a discrete r.v. is calculated as

$$\text{var}(X) = \sum_{x \in S_x} \{x - E(X)\}^2 f_X(x) \quad (12)$$

It measures the 'spread' of the r.v. X , i.e. the extent to which the distribution of X is dispersed around its mean μ_x .

It can also be calculated as

$$\text{var}(X) = E(X^2) - E\{X\}^2. \quad (13)$$

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The positive square root of the variance is called the **standard deviation (s.d.)** of X ,

$$\sigma_x = \sqrt{\text{var}(X)}$$

Properties of variance

- (i) $\text{var}(X) \geq 0$
with equality if and only if X is a constant.
- (ii) If $a, b \in \mathbb{R}$ are constants, then
 $\text{var}(a + bX) = b^2 \text{var}(X)$.

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3.4 Independence of r.v.'s

Definition

The discrete r.v.'s X, Y are said to be independent of each other iff

$$P(X = x \cap Y = y) = P(X = x) \times P(Y = y)$$

$$\forall x \in S_x, \forall y \in S_y.$$

Properties

If X, Y are independent r.v.'s, then

(i) $E(XY) = E(X)E(Y)$

(ii) $\text{var}(X + Y) = \text{var}(X) + \text{var}(Y)$ [Prove]

Counter **example** for property (ii):

If $X = Y$ (NOT independent), then

$$\begin{aligned}\text{var}(X + Y) &= \text{var}(2X) \\ &= 4\text{var}(X) \neq \text{var}(X) + \text{var}(Y)\end{aligned}$$