

HERIOT-WATT UNIVERSITY

<.SC. IN ACTUARIAL SCIENCE

Life Insurance Mathematics I

Tutorial 2 Solutions

1. (a) ${}_{10}p_{30:40}$ is the probability that (30) and (40) both survive 10 years: ${}_{10}p_{30} \cdot {}_{10}p_{40} = 0.97853$.
 - (b) $q_{30:40}$ is the probability that one or both of (30) and (40) die within one year: $1 - p_{30:40} = 0.001526$
 - (c) $\mu_{40:50}$ multiplied by a small time element dt is interpreted as the probability that (40) or (50) or both die within time dt : $\mu_{40} + \mu_{50} = 0.003274$
 - (d) ${}_{10}p_{[30]:[40]}$ is as for (a) but on a select basis: ${}_{10}p_{[30]} \cdot {}_{10}p_{[40]} = 0.97887$
 - (e) $q_{[30]:[40]}$ is as for (b) but on a select basis: $1 - p_{[30]:[40]} = 0.001264$
 - (f) $\mu_{[40]:[50]}$ is as for (c) but on a select basis: $\mu_{[40]} + \mu_{[50]} = 0.002293$
 - (g) $\mu_{[40]+1:[60]+1}$ as for (f) but on select basis for (41) and (61) both with select duration 1: $\mu_{[40]+1:[60]+1} = 0.008129$.
 - (h) ${}_3|q_{[30]+1:[40]+1}$ is the probability that one or both of (31) and (41), each with select duration of 1, will die within one year deferred for three years: ${}_3|q_{[30]+1:[40]+1} = 0.001976$
2. (a) The CDF of T_{max} , $P(T_{max} \leq t)$, denoted ${}_tq_{\overline{xy}}$, can be given as ${}_tq_{xt}q_y$. The density is therefore

$$\begin{aligned}
 f_{\overline{xy}}(t) &= \frac{d}{dt} {}_tq_{\overline{xy}} = \frac{d}{dt} {}_tq_{xt}q_y = \frac{d}{dt} (1 - {}_tp_x - {}_tp_y + {}_tp_{xt}p_y) \\
 &= {}_tp_x\mu_{x+t} + {}_tp_y\mu_{y+t} - {}_tp_{xt}p_y(\mu_{x+t} + \mu_{y+t}) \\
 &= {}_tp_x\mu_{x+t} + {}_tp_y\mu_{y+t} - {}_tp_{xy}\mu_{x+t:y+t}
 \end{aligned}$$

- (b) The density of T_{min} is ${}_tp_{xy}\mu_{x+t:y+t}$. Therefore its the expected value is given by $E[T_{min}] = \int_{t=0}^{\infty} t \cdot {}_tp_{xy}\mu_{x+t:y+t} dt$. Applying integration by parts, we let $u = t$ such that $\frac{du}{dt} = 1$ and we let $\frac{dv}{dt} = {}_tp_{xy}\mu_{x+t:y+t}$ such that $v = -{}_tp_{xy}$.

$$E[T_{min}] = -t \cdot {}_tp_{xy} \Big|_{t=0}^{t=\infty} - \int_{t=0}^{\infty} -{}_tp_{xy} dt = \int_{t=0}^{\infty} {}_tp_{xy} dt.$$

(c)

$$\begin{aligned}
\text{COV}(T_{min}, T_{max}) &= \text{E}[T_{min}T_{max}] - \text{E}[T_{min}] \cdot \text{E}[T_{max}] \\
&= \text{E}[T_x] \cdot \text{E}[T_y] - \overset{\circ}{e}_{xy} (\overset{\circ}{e}_x + \overset{\circ}{e}_y - \overset{\circ}{e}_{xy}) \\
&= \overset{\circ}{e}_x \overset{\circ}{e}_y - \overset{\circ}{e}_{xy} (\overset{\circ}{e}_x + \overset{\circ}{e}_y - \overset{\circ}{e}_{xy}) = (\overset{\circ}{e}_x - \overset{\circ}{e}_{xy}) (\overset{\circ}{e}_y - \overset{\circ}{e}_{xy}).
\end{aligned}$$

(d) i. The density of K_{min} is ${}_t|q_{xy}$ such that

$$\begin{aligned}
\text{E}[K_{min}] &= \sum_{k=0}^{k=\infty} k \cdot {}_k|q_{xy} = \sum_{k=0}^{k=\infty} k ({}_k p_{xy} - {}_{k+1} p_{xy}) \\
&= 0({}_0 p_{xy} - {}_1 p_{xy}) + 1({}_1 p_{xy} - {}_2 p_{xy}) + 2({}_2 p_{xy} - {}_3 p_{xy}) + \dots \\
&= {}_1 p_{xy} + {}_2 p_{xy} + {}_3 p_{xy} + \dots = \sum_{k=1}^{k=\infty} k p_{xy}.
\end{aligned}$$

ii.

$$\begin{aligned}
\overset{\circ}{e}_{xy} &= \int_{t=0}^{\infty} {}_t p_{xy} dt = \int_{t=0}^1 {}_t p_{xy} dt + \int_{t=1}^2 {}_t p_{xy} dt + \int_{t=2}^3 {}_t p_{xy} dt + \dots \\
&\approx 0.5({}_0 p_{xy} + {}_1 p_{xy}) + 0.5({}_1 p_{xy} + {}_2 p_{xy}) + 0.5({}_2 p_{xy} + {}_3 p_{xy}) + \dots \\
&= 0.5 + {}_1 p_{xy} + {}_2 p_{xy} + {}_3 p_{xy} + \dots = 0.5 + \sum_{k=1}^{k=\infty} k p_{xy} = 0.5 + e_{xy}.
\end{aligned}$$

(e) The ‘force of mortality’ associated with T_{max} can be defined as

$$\mu_{\overline{x+t:y+t}} = \frac{\overset{f}{\overline{x\overline{y}}}}{1 - \overline{F_{\overline{x\overline{y}}}}} = \frac{{}_t p_x \mu_{x+t} + {}_t p_y \mu_{y+t} - {}_t p_{xy} \mu_{x+t:y+t}}{{}_t p_{\overline{x\overline{y}}}}.$$

This way of defining a force is valid for any continuous random variable defining the time to a future event. For $t = 0$ we have $\mu_{\overline{x:y}} = 0$. This is surprising at first sight. However it is correct; the force $\mu_{\overline{x+t:y+t}}$ is defined in respect of the random variable $T_{max} = \max[T_x, T_y]$. If we choose another pair of ages $x' = x + s$ and $y' = y + s$ say, the force $\mu'_{\overline{x'+t:y'+t}}$ defined in respect of the random variable $T'_{max} = \max[T_{x'}, T_{y'}]$ is *not* the same as $\mu_{\overline{x+s+t:y+s+t}} = \mu_{\overline{x'+t:y'+t}}$. For the single life case, the random variables T_x and $T_{x'} = T_{x+s}$ were related by $P[T_{x'} \leq t] = P[T_x \leq s + t | T_x > s]$, and this was required in the proof that the forces defined in respect of T_x and $T_{x'}$ were equal at equal ages, namely $\mu'_{x'+t} = \mu_{x+s+t}$. There is no such relationship between $T_{max} = \max[T_x, T_y]$ and $T'_{max} = \max[T_{x'}, T_{y'}]$.

3. We have ${}_{10}p_x = \frac{l_{x+10:y}}{l_{x:y}} = 0.96$ and ${}_{10}p_y = \frac{l_{x:y+10}}{l_{x:y}} = 0.92$. Therefore the required probability is

$${}_{10}p_x (1 - {}_{10}p_y) + {}_{10}p_y (1 - {}_{10}p_x) = 0.1136.$$

4. (a) We note that this is the expected value of the random variable $\ddot{a}_{\overline{K_{min}+1}|}$. Therefore

$$\begin{aligned}\ddot{a}_{xy} &= \sum_{k=0}^{\infty} \ddot{a}_{\overline{k+1}|k} | q_{xy} = \sum_{k=0}^{\infty} \frac{1-v^{k+1}}{d} ({}_k p_{xy} - {}_{k+1} p_{xy}) \\ &= \frac{1}{d} \sum_{k=0}^{\infty} ({}_k p_{xy} - {}_{k+1} p_{xy} - v^{k+1} {}_k p_{xy} + v^{k+1} {}_{k+1} p_{xy}) \\ &= \frac{1}{d} \left(\sum_{k=0}^{\infty} {}_k p_{xy} - \sum_{k=0}^{\infty} {}_{k+1} p_{xy} - v \sum_{k=0}^{\infty} v^k {}_k p_{xy} + \sum_{k=0}^{\infty} v^{k+1} {}_{k+1} p_{xy} \right)\end{aligned}$$

But $\sum_{k=0}^{\infty} {}_{k+1} p_{xy} = \sum_{k=0}^{\infty} {}_k p_{xy} - 1$ and $\sum_{k=0}^{\infty} v^{k+1} {}_{k+1} p_{xy} = \sum_{k=0}^{\infty} v^k {}_k p_{xy} - 1$. Substituting gives

$$\ddot{a}_{xy} = \frac{1}{d} \left((1-v) \sum_{k=0}^{\infty} v^k {}_k p_{xy} \right) = \sum_{k=0}^{\infty} v^k {}_k p_{xy} \quad \text{since } d = 1-v.$$

- (b) $\ddot{a}_{xy:\overline{n}|}$ is the expected value of the random variable $\ddot{a}_{\overline{\min(K_{min}+1, n)}|}$:

$$\begin{aligned}\ddot{a}_{xy:\overline{n}|} &= \sum_{k=0}^{n-2} \ddot{a}_{\overline{k+1}|} \cdot k | q_{xy} + {}_{n-1} p_{xy} \cdot \ddot{a}_{\overline{n}|} = \sum_{k=0}^{n-2} \ddot{a}_{\overline{k+1}|} [{}_k p_{xy} - {}_{k+1} p_{xy}] + {}_{n-1} p_{xy} \cdot \ddot{a}_{\overline{n}|} \\ &= \sum_{k=0}^{n-2} \ddot{a}_{\overline{k+1}|} [({}_k p_x + {}_k p_y - {}_k p_{xy}) - ({}_{k+1} p_x + {}_{k+1} p_y - {}_{k+1} p_{xy})] \\ &\quad + ({}_{n-1} p_x + {}_{n-1} p_y - {}_{n-1} p_{xy}) \cdot \ddot{a}_{\overline{n}|} \\ &= \sum_{k=0}^{n-2} \ddot{a}_{\overline{k+1}|} [({}_k p_x - {}_{k+1} p_x) + ({}_k p_y - {}_{k+1} p_y) - ({}_k p_{xy} - {}_{k+1} p_{xy})] \\ &\quad + ({}_{n-1} p_x + {}_{n-1} p_y - {}_{n-1} p_{xy}) \cdot \ddot{a}_{\overline{n}|} \\ &= \left(\sum_{k=0}^{n-2} \ddot{a}_{\overline{k+1}|} \cdot k | q_x + {}_{n-1} p_x \cdot \ddot{a}_{\overline{n}|} \right) + \left(\sum_{k=0}^{n-2} \ddot{a}_{\overline{k+1}|} \cdot k | q_y + {}_{n-1} p_y \cdot \ddot{a}_{\overline{n}|} \right) \\ &\quad - \left(\sum_{k=0}^n \ddot{a}_{\overline{k+1}|} \cdot k | q_{xy} + {}_{n-1} p_{xy} \cdot \ddot{a}_{\overline{n}|} \right) = \ddot{a}_{x:\overline{n}|} + \ddot{a}_{y:\overline{n}|} - \ddot{a}_{xy:\overline{n}|}\end{aligned}$$

- (c) $A_{\overline{xy}}$ is the expected value of the random variable $v^{K_{max}+1}$.

$$\begin{aligned}A_{\overline{xy}} &= 1 - d\ddot{a}_{\overline{xy}} = 1 - d(\ddot{a}_x + \ddot{a}_y - \ddot{a}_{xy}) \\ &= (1 - d\ddot{a}_x) + (1 - d\ddot{a}_y) - (1 - d\ddot{a}_{xy}) = A_x + A_y - A_{xy}.\end{aligned}$$

5. We note that the expected value of the random variable $v^{K_{max}+1}$ is

$$A_{\overline{xy}} = \sum_{k=0}^{\infty} v^{k+1} \cdot k | q_{\overline{xy}} \quad \text{where } v = \frac{1}{1+i}.$$

The variance of $v^{K_{max}+1}$ is given by

$$\begin{aligned}\text{Var}[v^{K_{max}+1}] &= \text{E} \left[(v^{K_{max}+1})^2 \right] - (\text{E} [v^{K_{max}+1}])^2 \\ &= \sum_{k=0}^{\infty} (v^{k+1})^2 \cdot {}_k|q_{\overline{xy}} - (A_{\overline{xy}})^2 = \sum_{k=0}^{\infty} (v^2)^{k+1} \cdot {}_k|q_{\overline{xy}} - (A_{\overline{xy}})^2.\end{aligned}$$

For a rate of interest j we define $V = \frac{1}{1+j}$ and let $V = v^2$. This means that $j = i^2 + 2i$. Substituting in the above we get

$$\text{Var}[v^{K_{max}+1}] = \sum_{k=0}^{\infty} V^{k+1} \cdot {}_k|q_{\overline{xy}} - (A_{\overline{xy}})_{(\textcircled{i})}^2 = A_{\overline{xy}_{\textcircled{i} j=i^2+2i}} - (A_{\overline{xy}})_{(\textcircled{i})}^2.$$

6. (a) $\ddot{a}_{70:67} = 10.233$ (from tables).
- (b) $\ddot{a}_{70:67}^{(12)} \approx \ddot{a}_{70:67} - 0.458 = 9.775$.
- (c) $\ddot{a}_{70:67:\overline{10}|} = \ddot{a}_{70:67} - v^{10} {}_{10}p_{70}^m \cdot {}_{10}p_{67}^f \cdot \ddot{a}_{80:77} = 7.458$.
- (d) $\ddot{a}_{70:67:\overline{10}|}^{(12)} = (\ddot{a}_{70:67} - 0.458) - v^{10} {}_{10}p_{70}^m \cdot {}_{10}p_{67}^f \cdot (\ddot{a}_{80:77} - 0.458) = 7.204$.
- (e) $\ddot{a}_{\overline{70:67}} = \ddot{a}_{70}^m + \ddot{a}_{67}^f - \ddot{a}_{70:67} = 15.44$.
- (f) $\ddot{a}_{\overline{70:67}}^{(12)} = \ddot{a}_{\overline{70:67}} - 0.458 = 14.982$.
7. (a) $A_{\overline{xy}:\overline{n}|}$ is the EPV of an assurance of 1 payable at the end of the year in which the second of (x) and (y) dies, if that death occurs within n years, or 1 payable in n years exact if one or both (x) and (y) survive n years.

$$A_{\overline{xy}:\overline{n}|} = \text{E}[v^{\min[K_{\max}+1, n]}] = \text{E}[1 - d\ddot{a}_{\min[K_{\max}+1, n]}] = 1 - d\text{E}[\ddot{a}_{\min[K_{\max}+1, n]}] = 1 - d\ddot{a}_{\overline{xy}:\overline{n}|}.$$

- (b) $\bar{A}_{xy:\overline{n}|}$ is the EPV of an assurance of 1 payable immediately upon the first death of (x) or (y), if that death occurs within n years.

$$\bar{A}_{xy:\overline{n}|} = \text{E}[v^{\min[T_{\min}, n]}] = \text{E}[1 - \delta\bar{a}_{\min[T_{\min}, n]}] = 1 - \delta\bar{a}_{xy:\overline{n}|}.$$